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Feedback

Genetic and Morphometric Divergence of an Invasive Bird: The Introduced House Sparrow (*Passer domesticus*)in Brazil

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Abstract

Introduced species are interesting systems for the study of contemporary evolution in new environments because of their spatial and temporal scales. For this study we had three aims: (i) to determine how genetic diversity and genetic differentia‐ tion of introduced populations of the house sparrow (*Passer domesticus*) in Brazil varies with range expansion, (ii) to deter‐ mine how genetic diversity and differentiation in Brazil com‐ pares to ancestral European populations; and (iii) to deter‐ mine whether selection or genetic drift has been more influential on phenotypic divergence. We used six microsatellite markers to genotype six populations from Brazil and four populations from Europe. We found slightly reduced levels of genetic diversity in Brazilian compared to native European populations. However, among introduced populations of Brazil, we found no association between genetic diversity and time since introduction. Moreover, overall genetic differentia‐ tion among introduced populations was low indicating that the expansion took place from large populations in which genetic drift effects would likely have been weak. We found significant phenotypic divergence among sites in Brazil. Given the ab‐ sence of a spatial genetic pattern, divergent selection and not genetic drift seems to be the main force behind most of the phenotypic divergence encountered. Unravelling whether mi‐ croevolution (e.g., allele frequency change), phenotypic plas‐ ticity, or both mediated phenotypic divergence is challenging and will require experimental work (e.g., common garden ex‐ periments or breeding programs).

Introduction

Species invasions provide an opportunity to examine funda‐ mental questions in ecology and evolutionary biology, such as changes in geographical ranges, reproductive isolation and adaptation to novel environments, due to the large spatial and temporal scale of these "unplanned experiments" $[1]$. Bird introductions provide exceptionally good study opportunities because excellent historical records are often available, such as date of introduction, number of individuals released, num‐ ber of introductions and locality where individuals were released $[2]$, $[3]$. These data allow us to study evolution of species in new environments and over ecological time scales. More specifically, such instances generate data that can be used to examine how genetic diversity relates to range expan‐ sion $[4]$ – $[6]$ and the effects that selection and genetic drift may have on population divergence $[7]-[9]$ $[7]-[9]$. Most studies of non-native species have focused on *ecological* aspects of inva‐ sions, whereas *evolutionary* aspects have been less studied $[10]$, $[11]$. Therefore, incorporating the change in genetic and phenotypic properties due to evolution in the introduced envi‐ ronments may help to predict establishment success and im‐ pacts of non-native species $[12]$. For example, many introduced species only become invasive after a lag phase, which could be associated with the time that is necessary for evolu‐ tionary adjustments to take place $[13]$, $[14]$.

In general, the number of released individuals and introduc‐ tion events (propagule pressure) are associated with the suc‐ cess of establishment and spread of invasive species [\[15\]](#page-46-6).

These relationships are thought to exist because population size is tightly linked to demographic, environmental and ge-netic stochasticity [\[12\],](#page-46-3) [\[15\]](#page-46-6), [\[16\].](#page-46-7) Indeed, introduced populations tend to lose significant genetic diversity (i.e., allelic rich‐ ness and/or heterozygosity), because of founder events $[6]$. However, many invasive species show only modest reductions in genetic diversity $[17]$, which could be due to large propagule pressure, especially if propagules originated from different areas in the native range $[6]$, $[16]$, $[18]$. It is possible that for a significant decrease in genetic diversity to occur after an intro‐ duction event, a multiple step-wise colonization process (i.e., sequential founder events) may be necessary $[19]$, or in the case of multiple introductions, that gene flow in the intro‐ duced range be constrained $[18]$. In introduced birds, there is evidence for both loss of genetic diversity $[20]$ – $[25]$ and no change in genetic variability $[7]$, $[24]$, $[26]$, $[27]$. However, the loss of genetic variation in introduced bird populations is as‐ sociated with low propagule pressure and/or slow population growth rate after introduction $[23]$, $[28]$.

Studying genetic diversity and population structure across the range of a broadly distributed invasive species can help reveal the mechanisms that generate differentiation, as well as pro‐ vide insight into colonization dynamics $[4]$, $[12]$. For instance, the expansion of an invasive species can be a contiguous or non-contiguous process and can be accompanied by a large increase in the number of individuals, which together with the mode of dispersal will affect population genetic structure $[5]$, [\[29\]](#page-48-5), [\[30\].](#page-48-6) If dispersal (i.e., gene flow) between close populations is more frequent than between populations further

apart (moderate dispersal), an isolation by distance pattern should be expected. Conversely, no pattern of geographic ge‐ netic differentiation may occur in the introduced range if gene flow within the introduced range is strong relative to genetic drift, especially if the source propagule was genetically homo‐ geneous prior to the introduction (e.g. low propagule pres‐ sure). However, if gene flow is low among the expanding population fragments, then genetic drift will increase genetic dif‐ ferentiation and this process will be relatively independent of geographic distance. Over time, gene flow between adjacent population will form a pattern of isolation by distance but such an equilibrium between drift and gene flow might not de‐ velop during the limited time frame (e.g. a couple of hundred years) of most invasive species (see $Figure 1$ $Figure 1$ in $[5]$, $[31]$).

[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g001/) 1

Map of Brazil showing the different house sparrow populations sampled and the location of the release site.

To better understand rapid evolution and how invasive species adjust to novel environments, population genetic stud‐ ies should be combined with analyses of spatial phenotypic differentiation $[9]$. For example, some invasive species have the capacity to expand or shift their niches $[13]$, $[32]$, which is probably a response to novel selective pressures. If this is gen‐ erally the case, then a response in quantitative traits can be expected if there is enough genetic and/or quantitative varia‐ tion for selection to act upon $[17]$. In support, there are examples of introduced species that exhibit clinal patterns in mor‐ phology $[33]$, $[34]$, as might be expected with the above scenario. However, untangling whether selection or genetic drift

is responsible for phenotypic divergence is a complex task, and requires the initial step of comparing spatial genetic dif‐ ferentiation (F_{ST}) with spatial phenotypic differentiation (P_{ST}) . If P_{ST} is significantly larger or smaller than F_{ST} , then it is possible that the geographic variation in phenotypic traits were more likely shaped by selection rather than governed by ge-netic drift <u>[\[35\]](#page-49-3)–[37]</u>. P_{ST} is a rough estimate of Q_{ST} , the latter measures variation in quantitative traits by partitioning the variance related to additive genes between and within popula‐ tions [\[38\].](#page-49-5) However, attaining the necessary information for calculating Q_{ST} can be challenging because it requires the rearing of several populations in common garden conditions. Therefore, phenotypic measures have been used as a surro‐ gate, but one should be cautious to the possible caveats asso‐ ciated with the use of P_{ST} [\[39\]](#page-49-6).

In this study we used the house sparrow (*Passer domesticus*) to address three main questions. First, we compared genetic diversity and population structure between populations in the introduced range in Brazil and the native range in Europe. We expected populations from Brazil to be less diverse than the European populations. Second, we analyzed genetic differenti‐ ation among populations in Brazil to understand how the ex‐ pansion process in Brazil occurred. For example, if house sparrows in Brazil underwent sequential founder events dur‐ ing its expansion, one would expect a pattern of isolation by distance and populations in the expansion front to present re‐ duced genetic diversity. Our third goal was to compare spatial phenotypic differentiation (P_{ST}) with genetic differentiation (F_{ST}) . This method would allow us to evaluate whether phenotypic differences between populations were shaped by selec‐ tion or genetic drift. In principle, if P_{ST} equals F_{ST} , differentiation of morphological traits (assumed to be governed by addi‐ tive genetic variation) is probably the result of genetic drift. However, if P_{ST} is larger than F_{ST} it means that quantitative traits have diversified more than neutral genetic loci, which could be evidence of directional selection. Alternatively, if *P ST* is significantly smaller than F_{ST} , quantitative traits probably diversified less than neutral genetic loci, suggesting that these traits have been under the influence of stabilizing selection [\[34\]](#page-49-2)[–\[36\].](#page-49-7)

We chose the house sparrow as our model because of its remarkably broad distribution, predominantly determined by human introductions $[3]$, $[40]$. This distribution allows for multiple study replicates because genetic data from the intro‐ duced ranges of North America, Kenya, Australia and New Zealand $[24]$, $[26]$ already exist. Data on morphological divergence also exists from North America [\[34\],](#page-49-2) South America $[41]$ and New Zealand $[42]$.

Materials and Methods

Ethical Statement

This study was carried out in accordance with current laws of all countries where the study was performed and followed the recommendations of the Guidelines to the Use of Wild Birds in Research (Fair, J. E. Paul, and J. Jones, Eds 2010. Washigton,

D.C.: Ornithological Council). In Brazil approval by an ethical committee of the university is required only for captive ani‐ mals used in experiments. All other types of work with ani‐ mals are regulated by IBAMA - Instituto Brasileiro do Meio Ambiente e dos Recursos Naturais Renováveis, which substi‐ tutes the ethic committee in approving and evaluating all as‐ pects of projects that involve capture and handling of animals, taking morphological measurements, blood sampling and other procedures. Permits were obtained from IBAMA (179/2006-CGFAU; 123221 and 12322-2) for the purpose of this study and MR Lima's PhD Project and house sparrows were captured [with mist nets] at university campi and private homes *[in the cities of Brasília, Belém, Cáceres, Canoas, Niteroi* and Recife], with appropriate permissions of institutions and land owners. After being measured and sampled for blood and feathers, all birds were immediately released. Field meth‐ ods were carried out so that handling time and potential suf‐ fering of animals were minimized. House sparrows are consid‐ ered exotic birds in Brazil, and in the IUCN Red List the species has a Least Concern category. Data on European populations were obtained from a previous study [\[24\]](#page-48-1) with the permission of A Marzal and P Zehtindjiev. Blood samples from European house sparrows were obtained [with permission from A Marzal and P Zehtindjiev authors of] a previous PLoS ONE study $[43]$, which was approved by the Swedish Ethical Committee on Animal Experiment (reference M64-05).

Sampled Populations

Two hundred house sparrows were released in 1905 in Rio de Janeiro, Brazil $[44]$, and subsequent translocations and natural expansions of established populations have spread this species widely in Brazil, reaching the edges of the Brazilian Amazon in the city of Belém by 1978 $[44]$ – $[49]$. Six populations from Brazil were sampled and 15 individuals from each were genetically screened ([Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t001/) 1 and [Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g001/) 1). Data for four populations from Europe were obtained from a previous study [\(Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t001/) 1; see [Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g001/) 1 and [Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t001/) 1 in $[24]$). Data on the year that house sparrows arrived in the different sampled lo‐ cations in Brazil were obtained from the literature $[44]$ – $[49]$. It was not possible to sample house sparrows from Rio de Janeiro where they were initially released $[47]$, but we sampled house sparrows from Niteroi, which is 10 km from Rio de Janeiro. House sparrows in Brazil were caught using mist nets and blood was obtained from the brachial vein and conserved in 99% ethanol until DNA extraction.

Table 1

Sampled locations of Brazil (introduced) and Europe (native) with number of individuals genotyped (N) and captured (in parentheses) for which we have morphometric data, latitude and longitude in degrees, year that house sparrow arrived in the sampled location, mean number of alleles per loci (Na), allelic richness (Ar), private allelic richness (Par), observed heterozygosity (Ho), unbiased expected heterozygosity (UHe) and measure of departure form Hardy-Weinberg proportions (F_{IS}) .

Mean values in bold are significantly different (p<0.05).

1 – Using only the three matching loci.

Laboratory Procedures

Genomic DNA was extracted from blood samples using a stan‐ dard protocol with overnight digestion with proteinase K and subsequent phenol-chloroform extraction and alcohol precipi‐ tation [\[50\].](#page-51-0) Individuals were genotyped using six microsatel‐ lite loci (Pdoµ1, Pdoµ3, Pdoµ4, Pdoµ6, Pdo8 and Pdo9; $[51]$ -[\[53\]](#page-51-2), all of which were developed for house sparrows. Polymerase chain reaction (PCR) was performed in 10 µl reac‐ tions that contained 10 ng of template DNA, 5 µl of Qiagen multiplex master mix (contains pre-optimized concentrations of HotStarTaq DNA polymerase and \rm{MgCl}_{2} plus dNTPs and a PCR buffer especially developed for multiplex PCR), 1 pmol of each primer (forward primers were labelled with either 6- Fam or HEX) made up to 10 μ l with ddH₂0. For PCR conditions see [Information](#page-43-0) S1. We performed separate PCRs for the six loci.

PCR products of Pdou1, Pdou6 and Pdo8 were multiplexed and diluted 1∶100, while Pdoµ3, Pdoµ4 and Pdo9 were multi‐ plexed and diluted 1∶50. These multiplex combinations were chosen so that products had different dye labels and differed in range sizes. Labelled size standard MM1000 was mixed with multiplexed PCR products and electrophoresis was conducted in a capillary ABI3730XL sequencer (Applied Biosystems). Resulting data were analysed with GeneMapper 3.0 (Applied Biosystems) for fragment size determination.

Genotyping of house sparrow populations from Europe was done in ABI 377 (Applied Biosystems; see Schrey *et al.* [\[24\]](#page-48-1) for details), which does not use a capillary electrophoresis system. Therefore, 10 individuals sampled by Schrey *et al.* [\[24\]](#page-48-1) were genotyped together with individuals from Brazil to check for consistent allele scoring. For Pdoµ1 and Pdo9, we obtained a perfect match, whereas for Pdoµ3 there was a 2 bp differ‐ ence among the 10 individuals. Thus, we added 2 bp to the house sparrow sampled in Brazil to attain a perfect match with this locus as well. The remaining 3 loci did not match between studies, however, for Pdoµ4 and Pdoµ6, we obtained a perfect match for homozygosity and heterozygosity (i.e., indi‐ viduals that were homozygous and heterozygous in Schrey *et al.* [\[24\]](#page-48-1) were also homozygous and heterozygous in our analy‐ sis). We did not get a perfect match for Pdo8, thus, for the analyses below, genetic comparisons were done with and without the presence of Pdo8. Because results did not change when we excluded Pdo8 from the analysis, Pdo8 was maintained in the analysis. Unless otherwise stated, we only show results with Pdo8.

Genetic Diversity

For each of the six microsatellite loci and for each population, we tested for linkage disequilibrium (LE) and Hardy-Weinberg equilibrium (HWE) using FSTAT version 2.9.3 [\[54\].](#page-51-3) We observed no significant deviations from LE or HWE after correcting for multiple testing, except Pdoµ1 in Recife, Pdoµ6 in Spain and Pdoµ4 in Brasília and Italy that presented statistically sig-nificant heterozygote deficiency. We used Micro-Checker [\[55\]](#page-51-4)

to check for null alleles, large allele drop outs and stuttering. Indeed, Pdoµ1 in Recife had a high presence of null alleles (18%), as did Pdou6 in Spain (11%) and Pdou4 in Brasília (15%) and Italy (7%). However, when we pooled the data for analyses we saw no indication of true deviation from HWE ([Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t002/) 2). Because none of the loci consistently deviated from HWE or presented null alleles, it is likely that for the significant cases above, sampling error or infrequent cases of allelic dropout may have occurred. Moreover, at least for the popula‐ tions from Brazil, homozygote excess can be expected due to a founder effect.

Table 2

Polymorphic microsatellite loci used in genotyping house sparrow populations.

For each locus we list the number of alleles (Na), number of individ‐ uals types (N), observed heterozygosity (Ho) and expected heterozygosity.

To compare genetic diversity between the native range (Europe) and the introduced range (Brazil), we calculated al‐ lelic richness (Ar) and private allelic richness (Par) for each population using HP-Rare [\[56\]](#page-51-5) using all six loci, as well as just the three matching loci. When six loci were used, these calcu‐ lations were done separately for Brazilian and European pop‐ ulations. In the case of the Brazilian populations these esti‐ mates were calculated using a rarefication procedure with a minimum number of 28 alleles (smallest sample size=14 individuals), for each locus in each populations, while for European populations a minimum number of 16 alleles (small‐ est sample size=8 individuals) was used. For the three loci comparisons a minimum number of 16 alleles was used. Observed heterozygosity (Ho) and unbiased expected het‐ erozygosity (UHe) were calculated using GenAlEx version 6.1 [\[57\]](#page-51-6) and we used FSTAT version 2.9.3 [\[54\]](#page-51-3) to calculate num‐ ber of alleles (Na) and F_{IS} . We used non-parametric tests (Wilcox tests) to test for any differences in the genetic diver‐ sity estimators of the introduced and native house sparrow populations. We also compared these genetic diversity estima‐ tors with data from the literature of other house sparrow populations for studies that has used similar procedures to calculate these estimators $[24]$, $[58]$, $[59]$.

To test whether house sparrow populations in Brazil had ex‐ perienced a recent bottleneck, as might be expected if the ex‐ pansion process occurred via sequential founder effects or because of very small initial population size at the time of re‐ lease, we used BOTTLENECK version 1.2.02 [\[60\]](#page-52-2). The expected heterozygosity in BOTTLENECK was calculated under the Two-Phase Model (TPM) allowing for 95% single-step mutations and 5% multiple step mutations with a 12% variance for the multiple steps as recommended $[61]$. Significance of mismatch between expected and observed heterozygosity was inferred using the Wilcox test $[60]$. In addition, to test whether populations at the edge of expansion underwent sequential founder events, we subtracted the arrival year from 2012 (time since colonization) and used a Pearson correlation to test if there was a positive correlation between genetic diversity and time.

Population Structure

Genetic differentiation among the introduced populations of Brazil was determined by F_{ST} values, which were estimated according to Weir and Cockerham [\[62\]](#page-52-4) as implemented in FSTAT version 2.9.3 <u>[54]</u>. F_{ST} was estimated globally and between all pairs of introduced populations. Significance of global F_{ST} was evaluated by permutation of genotypes among samples and calculating 95% Confidence Intervals (C.I.) by bootstrapping over loci (number of permutations was set at 1000). Pairwise F_{ST} was tested to determine whether it was significantly different from zero by randomizing the genotypes, and a Bonferroni correction was used to control for type-I errors. We also calculated D_{est} defined by Jost <u>[\[63\]](#page-52-5)</u> because of the recent debates regarding F_{ST} calculations when using highly polymorphic markers such as microsatellites. D_{est} varies from zero, when there is no genetic differentiation between popula‐ tions, to one when populations are completely differentiated, and was calculated using the web-based resource SMOGD [\[64\]](#page-52-6) with 1000 bootstrap replicates and the harmonic mean of *D est* across loci. Moreover, we also used R_{ST} [\[65\]](#page-52-7) to infer population structure for populations from Brazil as implemented in R CALC [\[66\]](#page-53-0). This estimator is an analogue of F_{ST} , however, it uses variance in allele size (number of repeat units) between populations, because mutations in microsatellites involve the addition or subtraction of a small number of repeat units. We used R_{ST} because it is less sensitive to rare alleles than F_{ST} .

We tested for isolation by distance, which is the correlation between geographical distance (using log transformation) and the degree of genetic differentiation, using a Mantel test in Arlequin version 3.5.1.2. [<u>67]</u> for F_{ST} , while for $D_{es\;t}$ and R_{ST} we used the library "vegan" $[68]$ in R 2.14.0. We also calculated global and pairwise F_{ST} and D_{est} (as above) for the European populations. These calculations were repeated sepa‐ rately for the Brazilian and European populations because only three loci matched between the studies.

Phenotypic Data

Left tarsus, beak height, beak width and beak length of Brazilian sparrows were measured with a digital calliper (0.01 mm) and left wing, tail and body length were measured with a ruler (0.1 cm). Additionally, 770 feathers were plucked ran‐ domly from dorsal and breast areas (field procedures were conducted by MRL; samples sizes in [Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t001/) 1). There is no data on breeding period for house sparrow in Brazil, but all males had black beaks, which is indicative of breeding $[40]$, and we only found six individuals in active molt of remiges. Five feath‐ ers from each body region per individual were overlaid and taped to a black velvet substrate and feather colouration was measured using an Ocean Optics USB4000 spectrometer and a pulsed xenon light source (Ocean Optics PX-2; 220–800 nm range). All reflectance measurements were taken in relation to a WS-1SS white standard (Ocean optics, Dunedin, FL) and to the black velvet substrate (i.e., dark reference). We used a bi‐ furcated fiber-optic measurement probe, which was main‐

tained perpendicular to the feather surface at a fixed distance of 5 mm fixed to a probe block to eliminate external ambient light.

Spectrometric measurements were conducted with SpectraSuite software (Ocean optics) and three measure‐ ments, which consisted of 50 sequential spectra each, were taken from each sample at three random points by lifting the black block that contained the probe to ensure that a different part of the feather was being measured each time. Individual color was characterized by averaging the three spectra, which were interpolated to a step of 1 nm between 300 and 700 nm. We calculated brightness as the area under the spectra curve (i.e., value of zero meaning black and value of 100 meaning white) and UV-Chroma as the proportion of UV reflectance be‐ tween 300 and 400 nm.

Phenotypic divergence (P_{ST}) was used to infer the role of genetic drift and natural selection on the different morphological traits of house sparrow populations of Brazil by comparing it with F_{ST} . P_{ST} is similar to the Q_{ST} index, which measures quantitative trait differentiation, however, P_{ST} is influenced by environmental, non-additive genetic effects and by the interaction between the environment and genotype (see Merilä and Crnokrak $[36]$). Therefore, the use of P_{ST} as an approximation of Q_{ST} is usually not recommended [\[39\]](#page-49-6). However, to calculate Q_{ST} it is necessary to estimate the additive genetic variances, information that is obtained typically by rearing individuals from different populations in a common environment, which for several reasons, especially for vertebrate species, is not al‐

ways feasible. In the case of this study, the use of P_{ST} can be justified because Q_{ST} estimates are not available for our study populations and obtaining Q_{ST} would be very challenging (rearing of several house sparrow populations in commongarden conditions). Further, the morphometric traits being considered in this study are known to have substantial addi‐ tive genetic basis <u>[\[36\]](#page-49-7), [\[69\]](#page-53-3)</u>. Additionally, a P_{ST} – F_{ST} comparison can provide initial insights into the evolutionary process that has occurred during the expansion of the house sparrow in Brazil before further inquiries can be made. P_{ST} was estimated as:

$$
P_{ST} = \frac{\frac{c}{h^2} \sigma_B^2}{\frac{c}{h^2} \sigma_B^2 + \sigma_B^2}
$$

where σ_B^2 is the phenotypic variance between populations, σ_W^2 is the phenotypic variance within populations, and h^2 is the heritability (the proportion of the phenotypic variance attributed to additive genetic effects). The scalar *c* represents the proportion of the total variance that is claimed to occur be‐ cause of additive genetic effects across the populations. If pa‐ rameters c and h^2 are known for the populations being studied, then P_{ST} equals Q_{ST} [\[70\].](#page-53-4) However, estimation of c in the wild is very challenging and h^2 is population specific $[39]$. Because the c/h^2 ratio is critical to how well P_{ST} approximates Q_{ST} , one can use a sensitivity analysis, which varies this ratio, to infer the robustness of the approximation of Q_{ST} by P_{ST} [\[70\]](#page-53-4). According to this analysis, a null assumption would be to *ST*

consider $c/h^2 = 1$ (i.e., $c = h^2$), that is the proportion of phenotypic variance due to additive genetic effects is the same for both within and between population variance. If P_{ST} exceeds F_{ST} at this point it will also do so at any point where c >*h*² [\[70\]](#page-53-4). More important, however, is to evaluate if P_{ST} exceeds F_{ST} when c < h^2 (i.e., c/h^2 <1). The reason is that natural populations are probably under genotype-environmental interactions and/or divergent environmental effects and a low value of c/h^2 assumes a larger role of environmental effects in driving between population variance than within population variance (i.e, $c < h^2$). Therefore, the lower the critical c/h^2 ratio is (c/h^2 <1) when P_{ST} exceeds F_{ST} , the more likely it is that the trait is being shaped by selection $[70]$. Therefore, if there is evidence of between population variance deriving from addi‐ tive genetic effects, even in a scenario where environmental factors have a stronger role in determining phenotypic varia‐ tion, then phenotypic divergence will be the result of selection, as long as the trait is heritable [\[71\].](#page-53-5) *2* $_{ST}$ exceeds \bm{r}_{ST}

Variance components for estimating P_{ST} were obtained using analysis of variance where body length was entered as a co‐ variate. P_{ST} 95% C.I. were calculated, by considering P_{ST} to be normally distributed and using critical values of t, to test whether they overlapped with global F_{ST} value ±95% C.I. and thus whether P_{ST} values were different from F_{ST} . The critical c/h^2 ratio was obtained by graphically exploring P_{ST} and its 95% C.I. as a function of c/h^2 and by looking at the approximate value of c/h^2 where the lower 95% C.I. of P_{ST} meets the upper 95% C.I. of *F_{ST}* [<u>70</u>]. For example, a critical c/h^2 <0.1 means that in order for genetic drift to explain phenotypic di‐ *ST ST 2*

vergence, the required additive genetic effect across popula‐ tions would need to be less than 10% of the additive genetic effect within population. Therefore, this would be a very ro‐ bust inference that the traits are under selection and not ge-netic drift [\[70\]](#page-53-4).

We also conducted multivariate analysis of variance (MANOVA) to test if populationś centroids of trait means were significantly different from each other. MANOVA assumptions were checked before analysis, and three males and two fe‐ males were not included in the analysis because they were multivariate outliers, which was checked with the R package "mvoutlier" [\[72\]](#page-53-6).

Results

Impact of Introduction on Genetic Diversity

Brazilian house sparrow populations had significantly higher allelic richness ($W=23$, $p=0.02$) but significantly lower private allelic richness (W=24, p<0.01), and lower unbiased He (W= 23, $p=0.01$; see $Table 1$ $Table 1$ for mean and sd) than European house sparrow populations. When only three loci were used to calculate allelic richness and private allelic richness, Brazilian populations were significantly lower (respectively: W $=$ 24, p<0.01; W=24, p=0.01; see [Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t001/) 1 for mean and sd). However, there was no significant difference in the mean num‐ ber of alleles (W=8.5, p=0.52), observed heterozygosity (W= 6, $p=0.26$) and F_{IS} (W=8, $p=0.48$; see <u>[Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t001/) 1</u> for mean and

sd) between introduced (Brazil) and native (Europe) house sparrow populations. However, when genetic diversity is compared with other house sparrow studies, populations from Brazil did not present lower genetic diversity; for allelic rich‐ ness and private allelic richness they tended to present higher levels [\(Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g002/) 2). Although we are unable to test statistically because of differences in microsatellites and number of loci used between the different studies, house sparrow popula‐ tions in Brazil do not present a high loss of genetic variation. In support, none of the introduced populations from Brazil seemed to have experienced a significant bottleneck effect (lowest Wilcox one-tailed (heterozygosity excess) probability of 0.22). In the introduced range, there was no correlation be‐ tween time since colonization and any of the genetic diversity indices (Na: r=−0.02, df=4, p=0.97; Ar: r=−0.01, df=4, p=0.99; Par: $r=0.24$, $df=4$, $p=0.64$; Ho: $r=0.15$, $df=4$, $p=0.78$). Altogether, our results do not support a scenario of sequential bottlenecks during the house sparrow expansion in Brazil or a major loss of genetic diversity.

[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g002/) 2

Comparison of different genetic diversity estimators: (Na) number of alleles (A); (Ar) allelic richness (B); (Par) private allelic richness (C); and (He) expected heterozygosity (D) from different house sparrow populations.

For Europe –a and USA data from $[24]$; data for Finland from $[58]$; data for France from $[59]$; and data from Brazil and Europe –b where obtained from this study using all six loci [\(Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t001/) 1). Filled circles are introduced populations while open circles are native popula‐ tions. Not all estimators were available in all the studies.

Population Genetic Differentiation

Genetic differentiation among European house sparrow popu‐ lations was very low, both globally (F_{ST} among European population =0.019; 95% C.I: 0.010–0.031) and in pair-wise com‐ parisons (from 0.0043 to 0.0328; [Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t003/) 3). However, all pairwise F_{ST} values were significantly different for all European populations, except Italy and Spain. D_{est} values for the differ-ent European populations were high [\(Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t003/) 3), suggesting that genetic differentiation is present in Europe.

Table 3

Pairwise F_{ST} **values for house sparrow populations in Europe (lower diagonal), values in bold are significantly different from zero after Bonferroni correction (p≤0.0083) and harmonic** *D est***values (above the diagonal).**

For Brazil, genetic differentiation was also very low both glob‐ ally (F_{ST} among Brazilian population =0.028; 95% C.I: 0.016– 0.046) and between population pairs (from 0.0050 to 0.0695; [Table](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/table/pone-0053332-t004/) 4). However, two populations, Canoas and Niteroi, were significantly different from all other populations, and their pairwise F_{ST} value was highest among all pairwise values (0.0695). Canoas is in the South of Brazil, while Niteroi is less than 10 km from Rio de Janeiro, where the house sparrows were initially released (**Figure 1**). Belém and Recife were also significantly differentiated. The pairwise D_{est} values showed a similar pattern to the F_{ST} values (<u>Table 4</u>), ranging from 0.0161 to 0.2510 and were highly correlated with F_{ST} (Mantel r=0.79, p=0.013, 1000 randomisations); again, Niteroi and Canoas had the highest D_{est} value. Furthermore, when R_{ST} was used to infer genetic differentiation between house sparrow populations from Brazil, we found that global R_{ST} values were higher than zero $(R_{ST}$ among Brazilian populations = $0.033;$ 95% C.I.: $0.031–0.1111$) and pairwise R_{ST} values were similar to F_{ST} and D_{est} (<u>Table 5</u>) and highly correlated with F_{ST} (Mantel r=0.70, p=0.035). The results from the F_{ST} , D_{est} and R_{ST} analysis suggest slight genetic population differentiation in Brazil.

Table 4

Pairwise *F* **values for house sparrow populations in Brazil** *ST* **(lower diagonal), values in bold are significantly different from zero after Bonferroni correction (p≤0.0033) and harmonic** *D est* **values (above the diagonal).**

1 - City closest to place of initial introduction.

Table 5

1 - City closest to place of initial introduction.

We found no isolation by distance, as shown by the non-significant negative correlation between genetic differentiation and geographic distance (*F_{ST}*: Mantel r=−0.38, p=0.13, 1000 randomisations (<u>Figure 3A</u>); *D_{est}*: Mantel r=−0.10, p=0.59, 1000 randomisations (<u>[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g003/) 3B</u>); *R_{ST}*: Mantel r=0.03, p=0.49 ([Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g003/) 3C)).

[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g003/) 3

Scatterplots of F_{ST} **pairwise estimates [\[62\]](#page-52-4) calculated using FSTAT version 2.9.3 [\[54\]](#page-51-3) against geographical distance in km (log-transformed) for house sparrow populations of Brazil (A); pairwise harmonic mean** *D* **[\[63\]](#page-52-5) calculated using SMOGD** *est* **[\[64\]](#page-52-6) against geographic distance in km (log-transformed) for house sparrow populations of Brazil (B); and pairwise** *R ST* **calculated using R CALC [\[66\]](#page-53-0) against geographic distance in km (log-transformed) for house sparrow populations of Brazil (C).**

Phenotypic Differentiation in Brazil

Male and female morphologies differed, as shown by differ‐ ences in population centroids (Female: *Pillai trace* =2.246, df= 55, 245 F=3.633, p<0.001; Male: *Pillai trace* =1.897, df=55, 370, F=4.113, p<0.001). Comparisons of P_{ST} with F_{ST} show that similar traits in both males and females were usually shaped by selection (i.e., P_{ST} higher than F_{ST}) and not genetic drift, because lower 95% C.I. for P_{ST} were higher than the upper 95% C.I. for F_{ST} <u>[\(Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g004/) 4</u> and <u>[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g005/) 5</u>). Evidence for the

robustness of P_{ST} > F_{ST} varied among the traits but was exceptionally strong for plumage traits, which had critical c/h^2 lower than 0.10. Thus, the proportion of phenotypic variance across populations that is explained by additive genetic effects for plumage traits would need to be 10 times lower than the phe‐ notypic variation encountered within populations for these traits to be explained by genetic drift. For tarsus length the ad‐ ditive genetic effects would need to be 5 times lower, while for wing length it would be less than two times. However, we also found similar traits in both males and females were P_{ST} was either not higher than F_{ST} or when higher critical c/h^2 was usually between 0.5 and 1.2 ($Figure S1$ $Figure S1$ and $Figure S2$), indicating that these traits are probably shaped by genetic drift. Therefore, phenotypic differentiation is low for these traits and the inference of selection acting on these traits is less ro‐ bust than for the traits in $Figure 4$ $Figure 4$ and $Figure 5$. *2*

[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g004/) 4

Comparison of phenotypic differentiation $(P_{ST}$ – solid line) with **the upper 95% confidence interval (C.I.) for neutral genetic** differentiation (F_{ST} , solid red line), while the ratio c/h^2 was **varied from zero to 2.**

The dashed black lines represents the 95% C.I. for the P_{ST} calculations, while the dashed red line represent the null assumption that *c* $= h^2$. Results are for male traits that had critical c/h^2 (the value in which the lower 95% C.I. of P_{ST} is higher than the upper 95% C.I. of *F_{ST}*) lower than 0.5. For values with higher critical value see <u>[Fig.](#page-42-0) S1</u>.

[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g005/) 5

Comparison of phenotypic differentiation $(P_{ST}$ – solid line) with **the upper 95% confidence interval (C.I.) for neutral genetic** differentiation (F_{ST} , solid red line), while the ratio c/h^2 was **varied from zero to 2.**

The dashed black lines represents the 95% C.I. for the P_{ST} calculations, while the dashed red line represent the null assumption that *c* $= h^2$. Results are for female traits that had critical c/h^2 (the value in which the lower 95% C.I. of P_{ST} is higher than the upper 95% C.I. of *F_{ST}*) lower than 0.5. For values with higher critical value see [Fig.](#page-43-1) S2.

Discussion

The genetic variation in house sparrows from Brazil was only marginally lower compared to populations from the native range in Europe. We found no evidence for recent population bottlenecks or for the occurrence of sequential founder events during the range expansion process. We also found an absence of genetic structuring (or at most, weak structuring) among Brazilian populations, implying that expansion oc‐ curred with low influence of genetic drift and possibly high population growth. Moreover, we found that populations from Brazil differed morphologically from each other and that phe‐ notypic divergence (*P_{ST}*) was generally higher than expected from neutral genetic markers for similar traits in both males and females. However, our results must be interpreted with caution because of the small number of loci and populations used.

Genetic Diversity

Private allelic richness (Par), and unbiased expected heterozy‐ gosity (UHe) were lower in introduced Brazilian than native European house sparrow populations. When only three loci were considered, both allelic richness (Ar) and Par were lower for introduced Brazilian populations. These results are consistent with founder effects observed with other bird in‐ troductions $[20]$ – $[23]$, $[25]$. However, observed heterozygosity (Ho), number of alleles (Na), and inbreeding (F_{IS}) were not different from native European populations, and when six loci

were used, we found higher Ar for introduced Brazilian house sparrow populations. Therefore, house sparrow populations in Brazil do not appear to have lost much genetic variation. In support, when genetic diversity estimators from this study were compared with the literature, we found that estimators were not substantially different from what is found in popula‐ tions from the native range (**[Figure](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/figure/pone-0053332-g002/) 2**). Additionally, bottleneck signatures could not be detected for any Brazilian populations, although our small sample size may constrain our statistical power [\[60\].](#page-52-2) Moreover, time since colonization had no effect on genetic diversity, which suggests no occurrence of bottle‐ neck or sequential founder events during the house sparrow expansion in Brazil. Thus, it seems that house sparrows in Brazil did not go through a strong population bottleneck and/or that once introduced to Brazil, population size quickly increased, thus reducing the effect of genetic drift $[17]$, $[30]$.

House sparrow studies from other introduced ranges have shown mixed results regarding the amount of genetic diversity lost when compared with the native range. For example, intro‐ duced populations in Australia and New Zealand exhibit a re‐ duction in the number of alleles, but only the New Zealand populations had lower heterozygosity $[26]$. In North America, house sparrow populations had similar genetic diversity to na‐ tive European populations, and in Kenya, introduced popula‐ tions exhibit low levels of genetic diversity [\[24\].](#page-48-1) Differences in propagule pressure could explain the genetic diversity differ‐ ences found in these distinctive introduced ranges. If so, intro‐ duced ranges derived from low propagule pressure should ex‐ hibit significant losses in genetic diversity, while introduced

ranges from high propagule pressure should not present a re‐ duction in genetic diversity $[16]$, $[18]$. Although this pattern is maintained when we consider the North American introduc‐ tion, with a release of over 1000 individuals over several events [\[3\]](#page-45-2) with no reduction in genetic diversity, and the Australian and New Zealand introduction, which had over 300 individuals released over several events $[3]$ and showed significant genetic losses, the same cannot be said about the Brazilian introduction. In Brazil, 100 pairs were introduced [\[44\]](#page-50-3), therefore, significant losses in genetic diversity was expected, but substantial genetic loss was not found. It is possi‐ ble that initial propagule pressure for Brazil was higher than indicated by historical records. Problems with the accuracy in historical records have been shown in the past $[73]$ and care should be taken when using this kind of data to infer ecologi‐ cal processes. Another possibility is that rapid population growth could have occurred, which would have reduced the harmful effects associated with population bottleneck, allowing the retention of substantial genetic diversity $[30]$, especially if consecutive bottlenecks or founder effects did not oc‐ cur during the expansion [\[29\]](#page-48-5). Therefore, both differences in the introduction process and range expansions should influ‐ ence genetic diversity.

Population Genetic Differentiation in Brazil

We found low levels of genetic differentiation among house sparrow populations in Brazil. It seems that out of the six sam‐ pled populations only two, Canoas and Niteroi, are genetically different from all other populations. Canoas, which is in the

south of Brazil ($Figure 1$), may be influenced by other house sparrow expansions. For example, 20 pairs were introduced in Buenos Aires, Argentina, in 1872 $\boxed{3}$, and by 1888 house sparrows had already reached Uruguay $[47]$, both of which border the south of Brazil. Possibly, house sparrow popula‐ tions in the south of Brazil are an admixture of the two ex‐ panding populations, northward from Buenos Aires and southwards from Rio de Janeiro, which would explain why they are genetically different from the other populations in Brazil. However, Canoas presented the highest losses of genetic diversity $(Table 1)$ $(Table 1)$, which is not consistent with what would be expected for an admixed population. To test whether Canoas indeed has had a genetic influence from another ex‐ pansion front, populations from both Uruguay and Argentina would have to be sampled.

Although Niteroi is very close to Rio de Janeiro, the cities are separated by Guanabara Bay, a large body of water that the house sparrows would need to cross, or alternatively take a longer inland route between the two cities. Perhaps the popu‐ lation of Niteroi had a higher influence of genetic drift when compared to other populations, which could be the case if col‐ onization of Niteroi occurred before house sparrows could reach it via an inland route. Thus, Niteroi may not provide a good representation of the initial founding population of Rio de Janeiro. We also found no isolation by distance for popula‐ tions, and F_{ST} variation was not high, suggesting that the influence of genetic drift was low $[5]$, $[31]$. These results indicate that: (1) the source population was genetically homogeneous prior to the introduction (i.e., consistent with one introductory event); and (2) the expansion process probably occurred with high population growth and large propagule size from within the introduced range, which reduced the effects of genetic drift. In other introduced ranges, house sparrow populations also present weak genetic differentiation $[24]$, $[26]$, with the exception of Australia, in which populations are significantly more differentiated compared to New Zealand and Britain (most likely ancestral source population). Therefore, the low influence of genetic drift seems to be a common feature in the expansion process of house sparrows in introduced ranges. Moreover, house sparrows in the native range also present low levels of genetic differentiation $[24]$, $[58]$, $[59]$ and it is possible that the evolutionary history of house sparrows in the native range may have influenced the genetic diversity cap‐ tured during invasion $[74]$, i.e, house sparrow populations from the introduced range are simply reflecting the geograph‐ ical genetic structure of the native range. For example, the lack of genetic structure in the native range may result in low levels of population admixture in the introduced range.

A possible explanation for the low influence of genetic drift in house sparrow introduced ranges could be that there has not been enough time for genetic drift to take place because most introductions occurred around 1850 [\[3\]](#page-45-2). However, house sparrows are sedentary birds in their native range with natal dispersal distance of about 2 km $[40]$, and populations in the native range also present low genetic differentiation [\[24\]](#page-48-1), [\[58\]](#page-52-0), [\[59\].](#page-52-1) Therefore, gene flow may be comparatively high in this species. Data on dispersal distances in introduced ranges are available for North America, which show similarly short

dispersal distances $[40]$. If dispersal distance in the other introduced ranges is similar to that in the native range, which might be the case, and if colonization distance is also associ‐ ated with dispersal distance, then it is possible that the expan‐ sion of house sparrows in the introduced ranges (Brazil, North America and New Zealand) has been a contiguous process with high gene flow among the new founding popula‐ tions with high population growth.

Morphological Differentiation in Brazil

Morphological divergence was found among house sparrow populations in Brazil and P_{ST} - F_{ST} analysis indicates that divergence of most of the morphological traits was due to selection and not genetic drift, with the exception of wing length and most beak measurements.

House sparrow populations from other introduced ranges, such as North America [\[34\]](#page-49-2), New Zealand [\[42\]](#page-50-1) and Hawaii [\[75\]](#page-54-1), have also shown substantial morphological divergence. The latter study has also shown that morphological diver‐ gence was mainly due to selection and not genetic drift. Influence of genetic drift is higher in small populations [\[30\],](#page-48-6) but because house sparrows probably quickly expanded in the introduced ranges and, therefore, had large populations sizes, it can be expected that genetic drift did not play a sub‐ stantial role in the shaping of most morphological traits.

Although divergent selection (favouring of different pheno‐ types in different populations) may be driving morphological divergence of some of the traits, it is difficult to discern whether this pattern is a response to selection (microevolution) or simply a plastic response to the environment. In this study it is more challenging to evaluate this because we used P_{ST} instead of the more accurate Q_{ST} [\[39\]](#page-49-6) $\,$ Our P_{ST} estimates therefore cannot rule out environmental or parental (e.g. dif‐ ferences in parental care) effects on morphological traits. Nonetheless, without genetic differences between the intro‐ duced populations, it seems more plausible that phenotypic plasticity is driving morphological differentiation in some of the traits. However, if selection is indeed responsible for phe‐ notypic divergence, one can expect it to occur in a predictable manner, such as local adaptation to the abiotic environment [\[37\]](#page-49-4). It has already been shown that this could be the case for house sparrow populations of North America, where a posi‐ tive correlation was found between body size and latitude [34], [41]. However, no correlation was found between house sparrow morphological traits with latitude in South America $[41]$, which could be indicative of a lack of local adaptation and that genetic drift may be driving morphological diver‐ gence in South American populations. Our data show that traits related to body size, such as wing length, tarsus length and tail length did not present a robust critical c/h^2 when compared to other traits such as plumage coloration. Therefore, phenotypic differentiation of these traits may not have a very strong adaptive basis for house sparrow popula‐ tions in Brazil. A similar result was also found for native house sparrow populations from Finland, where only body mass

across populations seemed to be adaptive, while other traits (bill, wing and tarsus length) seemed to be shaped by genetic drift [\[71\]](#page-53-5).

However, we must interpret our results cautiously because of the small number of populations used in our $P_{ST} - F_{ST}$ analysis. Also, because two of the microsatellites had a high number of alleles, and therefore high-expected heterozygosity that can generate low levels of F_{ST} , it is possible that type-I errors of rejecting the null hypothesis P_{ST} = F_{ST} may have occurred [<u>9]</u>. However, our study is the first step in understanding the adap‐ tive potential of invasive populations of house sparrows in Brazil, and our initial data show that we may expect to find high plumage differentiation among populations of house sparrows in Brazil. Therefore, future efforts should explore why plumage may have a higher phenotypic differentiation when compared to other phenotypic traits.

Conclusions

Introduced house sparrow populations from Brazil lost some genetic variation relative to sparrows from the native range in Europe. However, it seems that the expansion process oc‐ curred in association with high population growth and possi‐ bly gene flow, thus enabling populations from Brazil to retain substantial genetic diversity with little genetic differentiation. However, our results need to be interpreted cautiously be‐ cause of the low number of markers and populations used. We found significant morphological variation among populations and, overall, morphological divergence was higher than

neutral genetic divergence suggesting the action of selection overriding the effect of genetic drift for many of the traits when F_{ST} was used. However, not all the traits presented P_{ST} > F_{ST} and traits related to body size (tarsus, wing and tail length) were less robust in the sensitive analysis then plumage traits. Using the P_{ST} – F_{ST} approach as an initial step allows us to infer that house sparrows should quickly respond to new selective factors they are exposed to in new areas, especially to factors affecting plumage coloration. In addition, future experimental studies should be able to determine if the morpho‐ logical divergence observed in Brazil is due to microevolution (changes in genotype frequency) or plastic phenotypic re‐ sponses to environmental conditions.

Supporting Information

Figure S1

Comparison of phenotypic differentiation (P_{ST} – solid line) **with the upper 95% confidence interval (C.I.) for neutral ge‐** ${\bf r}$ retic differentiation (F_{ST} , solid red line), while the ratio c/h^2 **was varied from zero to 2.** The dashed black lines the 95% C.I. for the P_{ST} calculations, while the dashed red line represent the null assumption that c = h^2 . Results are for male traits that had critical c/h^2 (the value in which the lower 95% C.I. of P_{ST} is higher than the upper 95% C.I. of F_{ST}) higher than 0.6.

(EPS)

Click here for [additional](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/bin/pone.0053332.s001.eps) data file.^(1.1M, eps)

Figure S2

Comparison of phenotypic differentiation (P_{ST} – solid line) **with the upper 95% confidence interval (C.I.) for neutral ge‐** ${\bf r}$ retic differentiation (F_{ST} , solid red line), while the ratio c/h^2 **was varied from zero to 2.** The dashed black lines represents the 95% C.I. for the P_{ST} calculations, while the dashed red line represent the null assumption that c = h^2 . Results are for female traits that had critical c/h^2 (the value in which the lower 95% C.I. of P_{ST} is higher than the upper 95% C.I. of F_{ST}) higher than 0.5.

(EPS)

Click here for [additional](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/bin/pone.0053332.s002.eps) data file.^(997K, eps)

Information S1

PCR Cycling conditions.

(DOCX)

Click here for [additional](https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3532305/bin/pone.0053332.s003.docx) data file.^(66K, docx)

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